

ON THE STRONG LAW OF LARGE NUMBERS FOR L-STATISTICS WITH DEPENDENT DATA

EVGENY BAKLANOV

ABSTRACT. The strong law of large numbers for linear combinations of functions of order statistics (L -statistics) based on weakly dependent random variables is proven. We also establish the Glivenko–Cantelli theorem for φ -mixing sequences of identically distributed random variables.

1. INTRODUCTION

Let X_1, X_2, \dots be a sequence of random variables with the common distribution function F . Let us consider the L -statistic

$$L_n = \frac{1}{n} \sum_{i=1}^n c_{ni} h(X_{n:i}), \quad (1)$$

where $X_{n:1} \leq \dots \leq X_{n:n}$ are the order statistics based on the sample $\{X_i, i \leq n\}$, h is a measurable function called a *kernel*, c_{ni} , $i = 1, \dots, n$, are some constants called *weights*.

The aim of this paper is to establish the strong law of large numbers (SLLN) for L -statistics (1) based on sequences of weakly dependent random variables. The similar problems were considered in the papers [1] and [2], where the SLLN was proved for aforementioned L -statistics based on stationary ergodic sequences. For example, in [2] the case of linear kernels ($h(x) = x$) and *asymptotic regular* weights was considered, i. e.

$$c_{ni} = n \int_{(i-1)/n}^{i/n} J_n(t) dt, \quad (2)$$

with J_n denoting an integrable function. In addition, the existence of a function J such that for all $t \in (0, 1)$

$$\int_0^t J_n(s) ds \rightarrow \int_0^t J(s) ds$$

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was imposed there. The statistics (1) with linear kernels and *regular* weights, i. e. $J_n \equiv J$ in (2), were considered in [1]. In the present paper we relax the regularity assumption on c_{ni} and, furthermore, consider the L -statistics (1) based on both stationary ergodic sequences and φ -mixing sequences. We also do not impose *monotonicity* of the kernel in (1). Note, that if h is a monotonic function, then the L -statistic (1) can be represented as a statistic

$$\frac{1}{n} \sum_{i=1}^n c_{ni} Y_{n:i},$$

based on a sample $\{Y_i = h(X_i), i \leq n\}$ (see [3] for more detail).

As an auxiliary result we obtain the Glivenko–Cantelli theorem for φ -mixing sequences.

2. NOTATIONS AND RESULTS

2.1. Assumptions and notations. We first introduce our main notations. Let $F^{-1}(t) = \inf\{x : F(x) \geq t\}$ be the quantile function corresponding to the distribution function F and let U_1, U_2, \dots be a sequence of uniformly distributed on $[0, 1]$ random variables. Due to the fact that joint distributions of random vectors $(X_{n:1}, \dots, X_{n:n})$ and $(F^{-1}(U_{n:1}), \dots, F^{-1}(U_{n:n}))$ coincide, we have that

$$L_n \stackrel{d}{=} \frac{1}{n} \sum_{i=1}^n c_{ni} H(U_{n:i}),$$

where $H(t) = h(F^{-1}(t))$, and $\stackrel{d}{=}$ denotes the equality in distribution. Let us consider a sequence of functions $c_n(t) = c_{ni}$, $t \in ((i-1)/n, i/n]$, $i = 1, \dots, n$, $c_n(0) = c_{n1}$. It is not difficult to see that in this case we have:

$$L_n = \int_0^1 c_n(t) H(G_n^{-1}(t)) dt,$$

where G_n^{-1} is the quantile function corresponding to the empirical distribution function G_n based on the sample $\{U_i, i \leq n\}$. We also introduce the following notation:

$$\mu_n = \int_0^1 c_n(t) H(t) dt,$$

$$C_n(q) = \begin{cases} n^{-1} \sum_{i=1}^n |c_{ni}|^q & \text{if } 1 \leq q < \infty, \\ \max_{i \leq n} |c_{ni}| & \text{if } q = \infty. \end{cases}$$

Further we will use the following conditions on the weights c_{ni} and the function H :

- (i) the function H is continuous on $[0, 1]$ and $\sup_{n \geq 1} C_n(1) < \infty$.
- (ii) $\mathbf{E}|h(X_1)|^p < \infty$ and $\sup_{n \geq 1} C_n(q) < \infty$ ($1 \leq p < \infty$, $1/p + 1/q = 1$).

Assumptions (i) and (ii) guarantee the existence of μ_n . We also note that $C_n(\infty) = \|c_n\|_\infty = \sup_{0 \leq t \leq 1} |c_n(t)|$ and $C_n(q) = \|c_n\|_q^q = \int_0^1 |c_n(t)|^q dt$ for $1 \leq q < \infty$.

2.2. SLLN for ergodic and stationary sequences. Let us formulate our main statement for stationary ergodic sequences.

Theorem 1. *Let $\{X_n, n \geq 1\}$ be a strictly stationary and ergodic sequence and let either (i) or (ii) hold. Then, as $n \rightarrow \infty$,*

$$L_n - \mu_n \rightarrow 0 \quad \text{a. s.} \quad (3)$$

REMARK. Let us consider the case of regular weights:

$$c_{ni} = n \int_{(i-1)/n}^{i/n} J(t) dt.$$

Then

$$L_n = \sum_{i=1}^n H(U_{n:i}) \int_{(i-1)/n}^{i/n} J(t) dt = \int_0^1 J(t) H(G_n^{-1}(t)) dt.$$

Hence, assuming $c_n(t) = J(t)$ in Theorem 1, we have

$$L_n \rightarrow \int_0^1 J(t) H(t) dt \quad \text{a. s.}$$

Also note that the convergence $\mu_n \rightarrow \mu$, $|\mu| < \infty$, yields that $L_n \rightarrow \mu$ a. s. In particular, if $c_n(t) \rightarrow c(t)$ uniformly in $t \in [0, 1]$, then $\mu_n \rightarrow \int_0^1 c(t) H(t) dt$.

Without the requirement that the coefficients c_{ni} are regular one can easily construct an example when the assumptions of Theorem 1 are satisfied, but the sequence $c_n(t)$ does not converges in any reasonable sense to a limit function. Let, for simplicity, $h(x) = x$ and let X_1 be uniformly distributed on $[0, 1]$. Set $c_{ni} = (i-1)\delta_n$ as $1 \leq i \leq k$ and $c_{ni} = (2k-i)\delta_n$ as $k+1 \leq i \leq 2k$, $k = k(n) = [n^{1/2}]$, $\delta_n = n^{-1/2}$. Thus, the function $c_n(t)$ is defined on the interval $[0, 2k/n]$. On the remaining part of $[0, 1]$ we extend $c_n(t)$ periodically with period $2k/n$: $c_n(t) = c_n(t - 2k/n)$, $2k/n \leq t \leq 1$ (see also [3, p. 138]). Note that $0 \leq c_n(t) \leq 1$. One can show that in this case $\mu_n \rightarrow 1/4$. In view of this fact we have that the assumptions of Theorem 1 are satisfied and, consequently,

$$L_n \rightarrow 1/4 \quad \text{a. s.}$$

2.3. SLLN for φ -mixing sequences. We will now formulate our main statement for mixing sequences. Let us define the mixing coefficients:

$$\varphi(n) = \sup_{k \geq 1} \sup \{ |\mathbf{P}(B|A) - \mathbf{P}(B)| : A \in \mathcal{F}_1^k, B \in \mathcal{F}_{k+n}^\infty, \mathbf{P}(A) > 0 \},$$

where \mathcal{F}_1^k and \mathcal{F}_{k+n}^∞ denote the σ -fields generated by $\{X_i, 1 \leq i \leq k\}$ and $\{X_i, i \geq k+n\}$ respectively. The sequence $\{X_i, i \geq 1\}$ is called φ -mixing (uniform mixing) if $\varphi(n) \rightarrow 0$ as $n \rightarrow \infty$.

Theorem 2. *Let $\{X_n, n \geq 1\}$ be a φ -mixing sequence of identically distributed random variables such that*

$$\sum_{n \geq 1} \varphi^{1/2}(2^n) < \infty, \quad (4)$$

and let any of the conditions (i) or (ii) hold. Then the statement (3) remains true.

The proof of Theorem 2 essentially uses the result of the Lemma 1 below. The statement (a) of Lemma 1 is the SLLN for φ -mixing sequences. The statement (b) is a Glivenko–Cantelli-type result for φ -mixing sequences and is of independent interest. We note that neither in Theorem 2 nor in Lemma 1 we do not assume the stationarity of the sequence $\{X_n\}$.

Lemma 1. *Let $\{X_n, n \geq 1\}$ be a φ -mixing sequence of identically distributed random variables such that the statement (4) holds. Then*

(a) *for any function f such that $\mathbf{E}|f(X_1)| < \infty$,*

$$\frac{1}{n} \sum_{i=1}^n f(X_i) \rightarrow \mathbf{E}f(X_1) \quad a. s. \quad (5)$$

(b)

$$\sup_{-\infty < x < \infty} |F_n(x) - F(x)| \rightarrow 0 \quad a. s., \quad (6)$$

where F_n is the empirical distribution function based on the sample $\{X_i, i \leq n\}$.

3. PROOFS

3.1. Proof of Theorem 1.

Lemma 2. *Let the function H be continuous on $[0, 1]$. Then*

$$\sup_{0 \leq t \leq 1} |H(G_n^{-1}(t)) - H(t)| \rightarrow 0 \quad a. s. \quad (7)$$

PROOF of Lemma 2. Using the equality

$$\sup_{0 \leq t \leq 1} |G_n^{-1}(t) - t| = \sup_{0 \leq t \leq 1} |G_n(t) - t|$$

(see, for example, [4, p. 95]) and the Glivenko–Cantelli theorem for stationary ergodic sequences, we get

$$\sup_{0 \leq t \leq 1} |G_n^{-1}(t) - t| \rightarrow 0 \quad a. s.,$$

i. e. $G_n^{-1}(t) \rightarrow t$ a. s. uniformly in $t \in [0, 1]$ as $n \rightarrow \infty$. Since the function H is uniformly continuous on the compact $[0, 1]$, it follows that $H(G_n^{-1}(t)) \rightarrow H(t)$ a. s. uniformly in $t \in [0, 1]$. This concludes the proof.

Let the condition (i) hold. Now, by Lemma 2,

$$\begin{aligned} |L_n - \mu_n| &\leq \int_0^1 |c_n(t)| |H(G_n^{-1}(t)) - H(t)| dt \\ &\leq C_n(1) \sup_{0 \leq t \leq 1} |H(G_n^{-1}(t)) - H(t)| \rightarrow 0 \quad a. s. \end{aligned}$$

Consequently, the proof of Theorem 1 for the first case is complete.

Lemma 3. *Let $\mathbf{E}|h(X_1)|^p < \infty$. Then*

$$\int_0^1 |H(G_n^{-1}(t)) - H(t)|^p dt \rightarrow 0 \quad a. s. \quad (8)$$

PROOF of Lemma 3. First note that the set of all continuous on the interval $[0, 1]$ functions is everywhere dense in $L_p[0, 1]$, $1 \leq p < \infty$. Therefore, for any $\varepsilon > 0$ and any function $f \in L_p[0, 1]$ there exists a continuous on $[0, 1]$ function f_ε such that $\int_0^1 |f(t) - f_\varepsilon(t)|^p dt < \varepsilon$. Since $\mathbf{E}|h(X_1)|^p = \int_0^1 |H(t)|^p dt < \infty$, this implies that there exists a continuous on $[0, 1]$ function H_ε such that

$$\int_0^1 |H(t) - H_\varepsilon(t)|^p dt < \varepsilon/2.$$

Further,

$$\begin{aligned} \int_0^1 |H(G_n^{-1}(t)) - H(t)|^p dt &\leq 3^{p-1} \int_0^1 |H(t) - H_\varepsilon(t)|^p dt \\ + 3^{p-1} \int_0^1 |H(G_n^{-1}(t)) - H_\varepsilon(G_n^{-1}(t))|^p dt &+ 3^{p-1} \int_0^1 |H_\varepsilon(G_n^{-1}(t)) - H_\varepsilon(t)|^p dt. \end{aligned} \quad (9)$$

From Lemma 2 it follows that $H_\varepsilon(G_n^{-1}(t)) \rightarrow H_\varepsilon(t)$ a. s. uniformly in t as $n \rightarrow \infty$. Hence, the last integral on the right hand side of (9) converges to zero a. s. as $n \rightarrow \infty$. Now let us consider the second integral. By ergodic theorem for stationary sequences,

$$\begin{aligned} &\int_0^1 |H(G_n^{-1}(t)) - H_\varepsilon(G_n^{-1}(t))|^p dt \\ &= \frac{1}{n} \sum_{i=1}^n |H(U_i) - H_\varepsilon(U_i)|^p \rightarrow_{\text{a. s.}} \mathbf{E}|H(U_1) - H_\varepsilon(U_1)|^p \\ &= \int_0^1 |H(t) - H_\varepsilon(t)|^p dt < \varepsilon/2. \end{aligned}$$

Consequently,

$$\limsup_{n \rightarrow \infty} \int_0^1 |H(G_n^{-1}(t)) - H(t)| dt < 3^{p-1} \varepsilon \quad \text{a. s.}$$

Since ε is arbitrary, we obtain (8).

Now let the assumption (ii) hold. Using Hölder's inequality, we get

$$|L_n - \mu_n| \leq C_n^{1/q}(q) \left(\int_0^1 |H(G_n^{-1}(t)) - H(t)|^p dt \right)^{1/p} \quad \text{for } p > 1,$$

and

$$|L_n - \mu_n| \leq C_n(\infty) \int_0^1 |H(G_n^{-1}(t)) - H(t)| dt \quad \text{for } p = 1.$$

The statement (3) follows from Lemma 3. This completes the proof of Theorem 1.

3.2. Proof of Theorem 2. We now prove Lemma 1. Note that for any measurable function f the sequence $\{f(X_n), n \geq 1\}$ has its φ -mixing coefficient bounded by the corresponding coefficient of the initial sequence, since for any measurable f the σ -field generated by $\{f(X_n), n \geq 1\}$ is contained in the σ -field generated by $\{X_n, n \geq 1\}$. Therefore, if the sequence $\{X_n, n \geq 1\}$ is φ -mixing, then so is the sequence $\{f(X_n), n \geq 1\}$. Hence, the condition (4) holds for mixing coefficients of the sequence $\{f(X_n), n \geq 1\}$. The statement (5) follows from the SLLN for φ -mixing sequences (see [5, p. 200]).

The statement (6) is an immediate corollary of (5) and classical Glivenko–Cantelli theorem.

The proof of Theorem 2 is similar to the proof of Theorem 1. Indeed, the statement (7) follows from the Glivenko–Cantelli theorem (6); using the SLLN (5), we get the statement (8). Thus the proof of Theorem 2 is complete.

REFERENCES

- [1] AARONSON, J., BURTON, R., DEHLING, H., GILAT, D., HILL, T. AND WEISS, B. (1996). Strong laws for L - and U -statistics. *Trans. Amer. Math. Soc.* **348** 2845–2866.
- [2] GILAT, D. AND HELMERS, R. (1997). On strong laws for generalized L -statistics with dependent data. *Comment. Math. Univ. Carolinae.* **38** 187–192.
- [3] BAKLANOV, E. A. AND BORISOV, I. S. (2003). Probability inequalities and limit theorems for generalized L -statistics. *Lithuanian Math. J.* **43** 125–140.
- [4] SHORACK, G. R. AND WELLNER, J. A. (1986). *Empirical processes with applications to statistics*. New York: John Wiley.
- [5] LIN, Z. Y. AND LU, C. R. (1996). *Limit theory for mixing dependent random variables*. Beijing: Kluwer.

DEPARTMENT OF MATHEMATICS
 NOVOSIBIRSK STATE UNIVERSITY
 PIROGOVA ST. 2
 NOVOSIBIRSK 630090, RUSSIA
E-mail address: baklanov@mmf.nsu.ru